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New Math Model for Independent-Model Triangulation

Applying the collinearity principle, the system may be used as a substitute for fully analytical aerial triangulation.

INTRODUCTION

T BIANGULATION by independent models is based on forming strips by joining independent models in planimetry, 4,9,10 or by spatial transformation, 13, 18, 19 or by performing the adjustment of relative orientation based on recorded y-parallaxes in the strip formation.11 In the adjustment phase, standard polynomial procedures are used in these methods. The polynomial adjustment procedures suffer from several shortcomings, 1,2,15 The mathematical model in this paper is based on utilizing two collinearity equations. The first one involves the model point, perspective center of the stereoplotter and the image point. The second one involves the object point on the ground, the exposure station and the image point.

It is believed that the use of this model should overcome the shortcoming of the polynomial adjustment. Also it will avoid:

- The problem in the scale determination which arises in the seven-parameters Anblock method,27,28
- The problem of assigning four parameters for planimetry and three others for height in the procedure described by v. d. Hout.¹⁰
- The problem of non-rigorous determination of the scale factor and the three shift parameters (which are determined not according to the least squares method) shown by Schut.¹⁸

At the same time, as shown in the paper, the suggested model provides complete statistical analyses for the observations and the parameters involved in the solution of this model.

DERIVATION OF THE MATHEMATICAL MODEL

The mathematical model is based on the

^o Presented at the 12th International Congress of Photogrammetry, Commission III, July-August 1972, Ottawa, Canada.

[†] For mathematical notations-see Appendix.

collinearity equation involving the terrain point, the exposure station and the image point. The image point coordinates are obtained using another collinearity equation involving the model point, the perspective center of the stereoplotter and the image point. Thus the mathematical formulation for an image point *i* in photo *k* transformed from model *j* is \dagger

$$\begin{bmatrix} \boldsymbol{V}_{X} \\ \boldsymbol{V}_{Y} \end{bmatrix}_{i,k} + \begin{bmatrix} X \\ Y \end{bmatrix}_{i,k} = f_{k} \begin{bmatrix} \overline{U}_{M} / \overline{W}_{M} \\ \overline{\Gamma}_{M} / \overline{W}_{M} \end{bmatrix}_{i,j}$$
$$= f_{k} \begin{bmatrix} \overline{U}_{G} / \overline{W}_{G} \\ \overline{\Gamma}_{G} / \overline{W}_{G} \end{bmatrix}_{i,j}$$
(1)

where

$$\begin{bmatrix} \overline{U}_{M} \\ \overline{V}_{M} \\ \overline{W}_{M} \end{bmatrix}_{i,k} = A_{R} \begin{bmatrix} U_{M_{i}} - U_{0} \\ V_{M_{i}} - V_{0} \\ W_{M_{i}} - W_{0} \end{bmatrix}$$
$$\begin{bmatrix} \overline{U}_{G} \\ \overline{V}_{G} \\ \overline{W}_{G} \end{bmatrix}_{i,k} = A_{E_{k}} \begin{bmatrix} U_{G_{i}} - U_{E_{k}} \\ V_{G_{i}} - V_{E_{k}} \\ W_{G_{i}} - W_{E_{k}} \end{bmatrix}.$$

THE DERIVED OBSERVATIONS

The derived observations are the image coordinates calculated using Formula 1 (for the left and right photos of a given stereomodel after performing a relative orientation) in the form:

$$Y_{1} = f' \cdot \begin{bmatrix} \overline{U}_{M}' / \overline{W}_{M}' \\ \overline{V}_{M}' / \overline{W}_{M}' \end{bmatrix}_{i,j}$$
(1A)

$$Y_2 = f^{\prime\prime} \cdot \begin{bmatrix} \overline{U}_M^{\prime\prime} / \overline{W}_M^{\prime\prime} \\ \overline{V}_M^{\prime\prime} / \overline{W}_M^{\prime\prime} \end{bmatrix}_{i,j}.$$
 (1B)

For above calculation, U'o, V'o, W'o and

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ABSTRACT: A new mathematical model for independent-model aerial triangulation is based on the collinearity principle. The validity of the model was assured using various sets of test data. The stereomodels are first relatively oriented; then the model coordinates are observed, and transformed to the image plane. These transformed coordinates are used as observations together with exposure station unknowns and the pass-point coordinates as parameters in a simultaneous least-squares solution. This mathematical model was searched for by using different geometrical approaches and various statistical models. The approaches to the geometrical models are based on the collinearity equation, conditions between unknown parameters for the exposure stations, and conditions between the transformed image coordinates. A suitable geometrical model was chosen from a comparative test of these approaches, using the collinearity equation as a basic equation in the adjustment. This model gives the best agreement between the observations and the parameters in the solution, uses a half-photograph as a basic unit in a simultaneous least-squares solution, and assigns for each basic unit three rotation parameters. The choice of the statistical model for this geometrical model is obtained by using different weight coefficient matrices for the transformed image coordinates as observations to obtain the best agreement between these observations and the parameters in the solution. For the specific statistical model that satisfies this condition, it is found that the transformed image coordinates can be treated as uncorrelated observations with equal weights. Inasmuch as the mathematical model was used for the first time in a simultaneous adjustment, the results were checked by statistical and geometrical analyses. Furthermore, the mathematical model is used in a comparison study with the existing analytical systems, using various test data on different stereoplotters and on one stereocomparator. From the results of these tests, it is concluded that: (1) independent-model aerial triangulation may be used as a substitute for analytical aerial triangulation; (2) the formulation of the adjustment of aerial triangulation must include the collinearity equation; and (3) mathematical correlation between the image coordinates may be neglected.

U''o, V''o, W''o, and the orthogonal matrices A'_{R} , A''_{R} should be used.

ADJUSTMENT PHASE

For any simultaneous adjustment using the image coordinates as observations, the mathematical model is

$$F(L^a, X_1^a, X_2^a) = 0.$$
 (2)

Linearizing Equation 2 leads to a linear observation equation having the form:

$$V = A_1 X_1 + A_2 X_2 + W.$$
 (3)

To find the values of X_i , X_2 and V of Equation 3 by the least-squares method, one needs the matrices (1) W, A_i , A_2 and (2) \overline{O}_{Y} .

Calculation of W, A_i , A_2 . Calculation of these matrices depends on the initial values for the parameters in the solution $(X_i^{\circ}, X_2^{\circ})$. Calculation of X_i° is accomplished using polynomial strip adjustment after strip formation is done using spatial transformation. For the calculation of X_2^o , space resection is performed either considering every whole photo or only half of it.

The Whole-Photo Case. The mathematical model in this case consists of Equations 2 and 3 in the form:

$$F(Y_{1}^{a}, Y_{2}^{a}, X_{2}^{a}) = 0$$
 (4A)

where
$$V = W + A_2 X_2 \theta$$
 (4B)

$$W = F(Y_1^{b}, Y_2^{b}, X_2^{0}).$$
 (4C)

In the space resection, the derived image coordinates Y_i and Y_2 are used as observations. Because the resection in photogrammetry consists of the determination of the six orientation parameters ($\omega, \varphi, \kappa, U_E, V_E$, W_E) from the positions of three or more image and object points, the values of $X_i^{\circ}, Y_i^{\circ}b$



FIG. 1. Convergence condition.

and Y_{2}^{b} are used for this purpose. The outcome of this phase is W, A_{1} and A_{2} , needed for the final least-squares adjustment of Equation 3.

The Half-Photo Case. A half-photo as a unit in space resection is treated because it is needed for mathematical model D (which is indicated later). The resection is performed by using either Y_{i}^{b} or Y_{2}^{b} as observations for solving X_{2}^{o} , which constitutes six parameters for each half-photograph. The mathematical formulation is:

For the left photo of model *j*,

 $F(Y_{I}^{a}, X_{o}^{a}) = 0$

and

$$V_i = A_2 X_2 + W_i \tag{5B}$$

where

$$W_1 = F(Y_1^{b}, X_2^{o}).$$
 (5C)

Similarly for the right photo of model *j*,

$$F(Y_2^a, X_2^a) = 0 \tag{6A}$$

and

$$V_2 = A_2 X_2 + W_2 = 0$$
 (6B)

where

$$W_2 = F(Y_2^b, X_2^o).$$
 (6C)

Calculation of $Q_{\mathbf{y}}$. Because Y_i and Y_2 are calculated using Formulas 1 and 2, and are functions of the original variables (the model coordinates M and the relative orientation elements X_R), then these formulas have to be linearized with respect to M and X_R . Furthermore, to study the radial weighing proposed by Hallert,⁸ and the effect of X_R on $Q_{\mathbf{y}}$, the following assumptions have been made.

$$(7A) \ \overline{Q}_{Y} = \sigma_{g}^{2} \cdot Q_{Y}$$

where

$$\sigma_{0}^{2} = \sigma_{01}^{2} + \sigma_{02}^{2}$$
 (7B)

(b) Linearizing, Formulas 1A and 1B are:

$$Y_{I} = B_{I} \cdot \Delta M + G_{I} \cdot \Delta X_{R}$$
(8A)

$$Y_{g} = B_{g} \cdot \Delta M + G_{g} \cdot \Delta X_{R}.$$
(8B)

(c) Applying error propagation to formulas 8A and 8B, one gets:

$$Q_{\mathbf{r}_{l}} = B_{l} \cdot Q_{\mathbf{M}} \cdot B_{l} t + K \cdot G_{l} \cdot Q_{\mathbf{R}} \cdot G_{l} t$$
(9A)

$$Q_{\mathbf{y}_2} = B_2 \cdot Q_M \cdot B_2 t + K \cdot G_2 \cdot Q_R \cdot G_2 t$$
(9B)

TABLE 1. DETERMINATION OF THE PERSPECTIVE CENTER AND RELATIVE ORIENTATION PARAMETERS

(5A)

Step	Determination of the Perspective Center	Relative Orientation Parameters
Observation	Projected grid points in X, Y plan	Y-parallax in the model measured with $\boldsymbol{\omega}$
Number of Points	9	6
Statistical Model	Equal weight and no correlation	Equal weight and no correlation
Estimate Required for	Approximate values for $U_0, V_0, W_0, \omega_0, \varphi_0, \kappa_0$	Approximate values for $\kappa', \kappa'', \varphi', \varphi'', \omega''$
Geometrical Model	Linear form of the collinearity equation	Linear form of the coplanarity equation
Output	Initial values for the zero setting of the instrument and the perspective center coordinates	The model coordinates, the relative orientation parameters and their estimated variance-covariance matrix
References	11, 15	5, 15
Remarks	No iteration is necessary (15)	Iterative procedure, the change is done only for the vector of misclosure

Case	Equation used for $Q_{\rm Y}$ and $Q_{\rm Y2}$	$Q_{\mathbf{Y}1\mathbf{Y}2}$	C_1	C_{2}	C_{3}	K
1	9A, 9B	9 C	4.1	0.017	0.000567	$\sigma_{\rm py}/\sigma_0$
2	9A, 9B		4.1	0.017	0.000567	1
3	$\begin{array}{llllllllllllllllllllllllllllllllllll$	_	4.1	0.017	0.000567	1
4	Diagonal term of Case 3		4.1	0.017	0.000567	1
5	$Q_{\rm v} = Q_{\rm r}$, 11A		4.1	0.017	0.000567	-
6	$\tilde{Q}_{v} = \tilde{Q}_{v}$, 11A		6.1	0.017	0.000567	-
7	$\tilde{Q}_{\rm v} = \tilde{U}_{\rm nit}$ Matrix	773	σ_{o}	0	0	

TABLE 2. CASES OF Q_y

$$Q_{\mathbf{Y}_1\mathbf{Y}_2} = K G_1 Q_R G_2^t. \tag{9C}$$

(d) Q_R can be obtained by utilizing the well known numerical procedure for relative orientation as shown in Table 1.

(e) To obtain Q_{M} , a space intersection can be used between the image and model coordinates:

 $O_{M} = (B^{t} O_{1}^{-1} B)$

where

$$Q_I = (1/\sigma_0^2) \cdot (S^2)$$
 (11A)

and

$$S = C_1 - C_2 r + C_3 r^2$$

(11B)

(10)

 C_i , C_2 and C_3 are chosen according to Hallert⁸ or similar assumptions; and they are introduced to take into account the radial weighing for the image coordinates. Their values are listed in Table 2.

(f) The previous assumptions have been made in various adjustment tests for deriving different patterns for $Q_{\rm r}$; and they are used as weight coefficient matrices in the least-squares solution for Equation 3. These assumptions are described briefly in Table 2. The main idea for these choices is to prove that:

- 1. It is not necessary to calculate Q_R and Q_M for the least-squares solution of the suggested mathematical model.
- 2. The transformed image coordinates $(Y_t \text{ and } Y_g)$, derived from the model coordinates after performing relative orientation, are mathematically uncorrelated. This proof is shown later.

CHECKING THE MATHEMATICAL DERIVATIONS

In order to check the mathematical derivations, some instrumental work was conducted on two stereoplotters (A8 and Kern PG2) and on one stereocomparator (STK1). The test data of this experiment are given in Table 3. Test number 1 was performed first to check the mathematical derivations given in the first section, and it was conducted in the following steps.

(a) Determination of the coordinates of the perspective center and the relative orientation parameters as shown in Table 1. The outcome of this stage is:

1. U'o, V'o, W'o and U"o, V"o, W"o

2. $\sigma_{02} = 4.23 \ \mu m$

3. X_R , Q_R and σ_{PY}

4. $U_{\rm M}$, $V_{\rm M}$ and $W_{\rm M}$.

(b) Using Equations 1A and 1B to obtain Y_1 and Y_2 .

(c) Determination of X_i^o by polynomial adjustment (Carlin²⁵) after strip formation.

(d) Calculation of different $Q_{\rm r}$ as outlined in Table 2.

(e) A simultaneous least-squares adjustment for Equation 2 using geometrical Approaches A,C,D,E and F (as shown in the next section).

The other test data are used when the suggested mathematical model (Approach D) is compared with other analytical systems in practice (as shown in a later section).

TABLE 3. TEST DATA

Test	Shediac Area New Brunswick	Swiss Test Block
Type of Photography	Wide Angle	Super Wide Angle
Scale	1/16,000	1/63,000 = 1/78,000
Principal Distance	152.07 mm	88.24 mm
Format	$230 \times 230 \text{ mm}$	$230 \times 230 \text{ mm}$
Number of Strips	1	2
Number of Models	6	8
Instrument Used	Wild A-8	K-PG2, STK1

Approach	X_2	Geometry		Basic Unit of Adjustment
A	6	$F(L^{a}, X_{a}, X_{a}) = 0$	(2A)	Photo
С	8	$F(L^{a}, X_{a}^{a}, X_{a}^{a}) = 0$	(2C)	Photo
D	9	$F(L^{a}, X_{a}^{a}, X_{a}^{a}) = 0$	(2D)	Half Photo
E	9	$F(L^{a}, X_{1}^{a}, X_{2}^{a}) = 0$	(2E)	Half Photo
F	.9	$ \begin{array}{l} F(X_{2R}^{a}) &= 0 \\ F(L^{a}, X_{1}^{a}, X_{2}^{a}) &= 0 \\ F(L^{a}) &= 0 \end{array} $	(2F)	Half Photo

TABLE 4A. THE GEOMETRICAL APPROACH USED

CHOICE OF THE BEST PARAMETERS

Different approaches to various geometrical models were tested; they are listed in Table 4A. The following are remarks on these approaches.

* Approaches A, C and D are based on Equation 2, the difference is the number of X_z per photo.

* Approach E used Equation 2 and the condition between unknown parameters X_2 , which implied that the orientation elements for the left and right halves of photo K should be equal. This condition makes Approach E identical with Approach A.

★ Approach F also used Equation 2 together with the condition between observations (Y_1, Y_2) , which implied that the space angle between any two identical pass points and the perspective center should equal zero, as shown in Figure 1. This condition occurs for a particular pass point in the triple lap area between two adjacent models. This condition is already implied in Approach D (and Equation 2) for the reason that in Approach D the collinearity condition is used for both rays specified in F. Hence, this condition is not an independent condition and it follows that geometrical models F and D should be identical. ★ The solutions of these geometrical approaches by a least-squares method are treated and developed by Maarek,¹⁵ following the principles given in References 20, 17, 24, 3.

★ The orientation parameters X_2 represent the six exterior orientations per photo in case of Approach A, the six exterior orientations per photo and two shift components for the inner orientation per photo for Approach C, and the nine orientation elements for Approaches D, E and F (they are three translation parameters for each photo (X_{2T}) and three rotational parameters (X_{2R}) for each half-photo). (Discussion of the latter point is given next.)

THE NECESSITY OF INTRODUCING ADDITIONAL THREE ROTATIONAL ELEMENTS

The necessity of introducing three rotational parameters while keeping the translation elements the same is justified by the following items.

■ The fact that the transformed image coordinates Y_{i}^{b} and Y_{2}^{b} (which are calculated by Equations 1A and 1B) for a particular intermediate photo in a certain strip come from different relatively oriented models.

 \blacksquare Y_{a}^{b} and Y_{a}^{b} are already rotated by the

		x_{2T}^{o} (in ft.).			xo2R Grads	
Photo	U_{E}^{o}	V_{E}^{o}	W^{o_E}	ω_{E}^{O}	φ^{o}_{E}	κo_E
1	1489597.8	900103.7	7974.80	0.8821	1616	199.8237
2R	1494356.1	900090.3	7987.20	0.8223	1784	199.8276
2L	1494351.7	900089.9	7987.40	-0.2525	-0.0322	199.9378
3R	1499488.7	900084.7	7989.13	-0.2476	-0.0139	199.9423
3L	1499487.5	900088.4	7987.80	-0.0352	0.1281	200.2045
4R	1504391.8	900102.5	7977.98	-0.0369	0.1278	200.2051
4L	1504390.9	900105.71	7974.96	0.4222	-0.1506	199.8838
5R	1509025.3	900097.8	7986.68	0.4160	-0.1618	199.8834
5L	1509030.1	900101.2	7986.80	0.3859	-0.2128	199.9089
6R	1513684.6	900057.3	8001.50	0.4169	-0.2130	199.3976
6L	1513685.0	900057.0	7999.10	0.5607	-0.0219	199,2506
7	1518305.7	900003.9	8000.00	0.5453	-0.0177	199.2515

TABLE 4B. APPROXIMATE PARAMETERS FOR X₂.

angular parameters of the relative orientation X_{R} .

The positions of these Y_i^b and Y_2^b are not related to a fixed coordinate system as in the case of a normal analytical triangulation measured on a monocomparator.

■ A study of the values of X_2^{o} , after space resection using the half-photo (Table 4B), shows that the differences between the translation elements for the right and the left photo do not exceed 2.5 feet (deviation from the mean). This is in the order of 30 μm at image scale, and in the order of the standard deviations of the perspective center of the stereoplotter. Furthermore, the vector of misclosure W in Equation 3 is always calculated by returning to the original Equation 2. Thus it is only necessary to introduce X_{zT} as three parameters per photo.

■ On the other hand, from the same Table 4B, the range of change between the sets of X_{2R}^{o} for the right and the left halves of the photos has reached the value of one grad. Thus, it is essential to introduce three rotational parameters for each half-photo. It is very essential to introduce this change because it does not correspond with the change (which does not exceed 0.0005 grad) of the rotational elements of the perspective center of the projectors of the stereoplotter.

THE CHOICE OF THE SUGGESTED MODEL Analyses of the different geometrical ap-

TABLE 4C. RESULTS OF USING THE GEOMETRICAL APPROACHES

Approach	Statistical Case	$\sigma_o^2(mm^2)$	df
Α	7	0.932	137
A	2	0.790	137
E	2	0.797	137
C	2	0.798	123
D	2	0.000349	122
F	7	0.000150	122

proaches using different Q_r (Table 2), were conducted to clarify two points: (1) the necessity of introducing three rotational parameters for each half of photo, and the best geometrical approach, and (2) the choice of weight coefficient matrix for the suggested model.

With respect to the first point, the results obtained are listed in Table 4C. From these results, it is obvious that the unbiased geometrical model corresponds to Approach D because its estimated variance of unit weight δ_{ρ}^2 is the smallest.

To study the second point, the different weight coefficients given in Table 2 are used, and the results obtained are shown in Table 5. Comparing the estimated variance $\hat{\sigma}_{a}^{2}$ of unit weight of each case with each other, and checking $\hat{\sigma}_{a}^{2}$ of each case with the tabulated value of the Fisher test, it is found that the unbiased $\hat{\sigma}_{a}^{2}$ corresponds with the statistical model Case 7 (Table 5). This leads to the conclusion that the transformed image coordinates behave as uncorrelated observations and with equal weights.

STATISTICAL VALIDITY OF THE RESULTS

Statistical tests have been applied for testing the behaviour of V and it is found that the following characteristics are valid.

▲ With respect to the estimated mean of V(U); it is found that U is an unbiased estimate by checking it against the theoretical mean U_t (U_t equals zero). The statistical test used in this case is the normal distribution test.

▲ With respect to the shape and the density of V° , V was plotted in histogram and it was found that it belongs to normal distribution with modulated structure Romanowski.^{22,23} The statistical test applied in this case was χ^2 -test.

 \blacktriangle The number of rejected observations were found as listed in Table 6.

TABLE 5. CHOICE OF THE WEIGHT COEFFICIENT MAT	FABLE	5. CHOICE OF TI	IE WEIGHT	COEFFICIENT	MATRIX
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		517+0	2	
Statistical Case	$mm \times 10^{-2}$	$mm^2 \times 10^{-3}$	$\sigma_0^2 \mu m^2$	Remarks
1	8.89	0.139	349.62	
2	8.60	0.130	349.62	
3	8.52	0.127	349.62	
4	8.52	0.127	349.62	
5	-1.65	-0.234	314.15	
6	8.21	0.222	290.64	
7	-0.079	-0.687	150.36	n = 304
7'	-0.095	-0.224	82.23	n = 298

^o See also "Modulated Normal Distribution and Photogrammetric Measurements," by the author, Photogrammetric Engineering, 39:8, 1973.

TABLE 6.	NUMBER	OF	REJECTI	ED OBSERVATIONS
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Test	No. of Observations	Rejected Observations
Grid Measurement	630	2
Test 1	304	6
Test 2	406	8

These rejected observations were relatively few, which agrees with the ideas of Rosenfield.²¹ Also these residuals have biased character and gross magnitude and they do not belong to the normal distribution; furthermore their magnitude is larger than 4 \mathfrak{F}_{q} .

▲ With respect to the estimated $\hat{\sigma}_o^2$, this is usually carried out by the use of x^2 -test as follows:

$$(1 - \alpha) = P \left[\frac{\dot{\sigma}_0^2 \cdot df}{\chi_{df, \alpha/2^2}} < \sigma_0^2 < \frac{\dot{\sigma}_0^2 \cdot df}{\chi_{df, (1 - \alpha/2)^2}} \right]$$

= 0.95. (12A)

The necessary statistical quantities for Equation 12A are given in Table 7.

 $\hat{\sigma}_o^2$ is obtained from Table 5, σ_o^2 is calculated using Equation 7B. Substituting in Equation 12A, one gets:

$$63.83 < 66.89 < 110.88$$
 (12B)

This means that $\hat{\sigma}_{o}^{2}$ is an unbiased estimated variance of unit weight.

▲ Finally, it is of great interest on concluding the statistical analyses for the suggested model, to check the power of the x^2 test on accepting ϑ_a^2 to represent the sample variance; in other words, one likes to know the probability of error Type II (B) when the x^2 -test is used to test the hypothesis concerning the variance ϑ_a^2 . Such a test is shown in Hamilton⁷ (page 83). Using this test, it is found the B = 60 percent.

Comparison of the Succested Mathematical Model with Other Triangulation Systems

In the following the obtained results from the two test data given in Table 3 were compared with the other triangulation sys-

TABLE 7. STATISTICAL QUANTITIES FOR Equation 12A

$\hat{\sigma}_0^{}2$	df	σ_0^2	χ ² 116,0.025	χ ² ,116,0.975	α
82.23	116	66.89	87.87	150.13	0.05

TABLE 8A. COMPARISON OF \overline{S}_{U} , \overline{S}_{V} , \overline{S}_{W} with RMSE from Anderson⁶, Values in Meters (*c* Represents Scale Factor)

Estimated Standard Deviation for X ₁	S _v	Sv	S_w	Remarks
Calculated	0.212	0.179	0.352	
$S \times c$	0.870	0.740	1.450	c = 4.11
RMSE	\overline{S}_{u}	\bar{S}_v	\overline{S}_{w}	(6)
Minimum	0.570	0.520	0.690	
Maximum	1.400	1.700	11.900	

tems in practice. Such a comparison is difficult due to the lack of information about an equivalent value to $\hat{\sigma}_{\rho}^2$ in the published aerial triangulation adjustments by other organizations. The limited size of the test materials used, the distribution of the ground control, bridging distance, type of photography and scale of photography question the validity of such comparison. Furthermore, in photogrammetric practice, RMSE (Anderson⁶) constitutes the normal criterion for comparison, for which either a test area is required¹² or simulated blocks⁶ must be used.

Comparison of the values of S_v , S_v and S_w by the results given by Anderson⁶ are shown in Table 8A where S_v , S_v and S_w are calculated using the equations,

$$S_U = \sqrt{(\Sigma \widehat{S}_{U_i}/n_u)}$$

$$S_V = \sqrt{(\Sigma \widehat{S}_{V_i}/n_v)}$$

$$S_W = \sqrt{(\Sigma \widehat{S}_{W_i}/n_w)}$$

$$S_P = \sqrt{(S_U^2 + S_v^2)}$$
(17)

where \hat{S}_v , \hat{S}_v , \hat{S}_w are obtained using the diagonal elements of the variance-covariance matrix of the adjusted pass point coordinates (E_{x_1}) after the least-squares adjustment of equation 2; n_v , N_v , n_w are the number of the coordinates of the pass points in X, Y, Z directions respectively. Another comparison is done between the estimated variance of unit

TABLE 8B. COMPARISON OF $\hat{\sigma}_0$, WITH $\sigma_{XY}(\sigma_X, \sigma_Y \text{ FROM JAKSIC}^{12}, \text{ p. 400, 401})$

	Computation		
	Unrefined Coordinates	Refined Coordinates	Remarks
σ	± 8.7	\pm 7.6	
σ.	\pm 7.4	\pm 6.2	(12)
σ _{vv}	± 11.4	± 9.8	
ô,	± 12.3	± 12.3	n = 304
Ū.	± 9.1	± 9.1	n = 298

	Simultaneous AMAREO (25)	Mathematical Model D	Polynomials (29)	Units
\$ _n	2.63	3.05	3.72	meters
S _v	2.87	3.26	3.81	meters
\overline{S}_{w}	2.93	3.76	4.18	meters
number of iterat	ions 2	1	15	

TABLE 9. RMSE OF SIMULTANEOUS, MATHEMATICAL MODEL D AND POLYNOMIAL ADJUSTMENTS (USING STEREOCOMPARATOR DATA; $n_c = 18$ Points

TABLE 10. RMSE OF MATHEMATICAL MODEL D AND POLYNOMIAL ADJUSTMENT (USING KERN PG2 DATA), $n_c = 18$ Points

	Polynomials	Mathema	tical Model D	Units
\$ _v	4.04	3.50	3.40	meters
\$v	4.00	3.82	3.23	meters
S _P	5.68	5.26	4.70	meters
S _w	3.68	4.03	3.16	meters
number of iteration	ons 12	1	2	motors

weight $\hat{\sigma}_o^2$ and the similar value given by Jaksic¹² shown in Table 8B.

The suggested mathematical model was used by Okuwa¹⁶ (in the case of the second test) to prove its validity; the results obtained are given in Tables 9, 10, 11. Table 9 shows the results using the suggested model, polynomial adjustment by Schut²⁹ and simultaneous adjustment described by U.S.C.G. and modified by Carlin.²⁵ The basis of the comparison in Table 9 is RMSE, using 18 check points. Table 10 shows a comparison between the suggested model and polynomial adjustment.²⁹ The instrument used in the comparison given in Table 9 was stereocomparator STK1 whereas for Table 10 it was Kern PG2.

As they were available, a comparison was made between the RMSE in planimetry and height for Kern PG2 (Table 10) and the values of S_P and S_W , calculated from Equation 17, shown in Table 11. This comparison shows that it is practically accurate to consider S_P and S_W as representative for RMSE in planimetry and height respectively, despite the fact that the statistical basis for such comparison is not so rigorous.

CONCLUSIONS

Considering the results achieved if the

test data are subjected to the suggested mathematical model, the following conclusions can be drawn.

C It is obvious from Tables 4A, 4B, and 4C that the best parameters for the solution came from mathematical model *D*, which assigned to each half-photo three rotations different from the other half.

C The derived observations, computed from model coordinates and relative orientation elements in a least-squares solution, behave as observations of equal weight and free from mathematical correlations. This becomes clear from Table 5 and from different choices of weight coefficient cases.

C The variances of the adjusted parameters of the coordinates of the pass points were checked and they were found to be satisfactory for the use in large scale mapping. This is obvious from the values of S_v , S_v and S_w , which are given in Table 8A.

C Comparison of the values of S_v , S_v and S_w by the results shown in (Anderson,⁶ Jaksic¹²); and the comparison of the suggested system with other systems in practice given in Tables 8A, 8B, 9 and 10 prove that independent-model aerial triangulation can be used as a substitute for analytical systems, providing that the suggested mathematical model is used in the adjustment.

TABLE 11. COMPARISON OF S_p and S_w with \overline{S}_p , \overline{S}_w (Values of \overline{S}_p , \overline{S}_w from Table 10)

			1000 Hollo			
Number of Iterations	S _v	S _v	Sp	\overline{S}_p/S_p	Sw	\overline{S}_w/S_w
1	3.22	4.05	5.26	1.00	5.25	0.74
2	2.87	3.56	4.62	1.02	3.66	0.86

C The applied statistical analyses prove that the suggested mathematical model is an unbiased mathematical model and applies a rigorous least-squares solution.

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APPENDIX—MATHEMATICAL INOTATIONS			raphy. It is assumed to be $49 \ \mu m \approx$ (8).		
i Image i Mode k Photo	e point. l number. number.	$\sigma_{02}^{}2$	A priori variance of unit weight related to the measuring instru- ment. σ_{o2} is obtained by grid meas- urement (Table 1).		
SUBSCRIPTS		S	Standard error of image coordi- nates.		
M Model coordinates. G Ground coordinates. T Translation elements. B Botational elements		$\sigma_{PY}^{}2$	The estimated variance of unit weight for the <i>y</i> -parallax measurements.		
SUPERSCRIPTS		σ_o^2	The estimated variance of unit		
t Trans a Adjus	pose. ted quantities.	_	tained from a least-squares adjust- ment.		
b Obser o Appro	ved quantities. ximate quantities. 	Q _y	Variance-covariance matrix for the image coordinates.		
" Right photo.		RMSE	Root mean squares error of the dis- crepancies on the check points.		
		S_{U}, S_{V}, S_{W}	RMSE in X , Y , Z directions.		
X, Y	Image coordinates in X , Y directions.	S_v, S_v, S_w	The unbiased estimated standard deviations for the pass point co- ordinates in X, Y, Z directions.		
V_x, V_y	Residuals of the image coordinates.	r	The radial distance of the image		
U_M, V_M, W_M	Model coordinates in X , Y , Z directions, respectively.	n	point. number of observations.		
f	Principal distance.	n_c	number of check points.		
Uo, Vo, Wo	The perspective center coordinates in the stereoplotter.	df	degree of freedom.		
U_E , V_E , W_E	U_E, V_E, W_E The exposure station coordinates for photo k.		EQUATIONS USED IN THE PAPER $(Vx, Xy)t = V_1t$		
$U_{\rm G}, V_{\rm G}, W_{\rm G}$	The ground coordinates in X , Y , Z directions, respectively.	$Z (Vx, Vy)^t = V_2^t$			
A_R	Orthogonal matrix, relates the im- age coordinates with model coor- dinates.	$(x, y)^{t} = Y_{1}^{t}$ $(x, y)^{t} = Y_{2}^{t}$ $(U_{M2} V_{M1}, W_{M2})^{t} = M^{t}$			
A_E	Orthogonal matrix, relates the im- age coordinates with ground co- ordinates.	$(U_G, V_G, W_G)t = X_1t$ $(U_E, V_E, W_E)_Kt = X_2t$			
μm	n micrometers.		$(\kappa, \omega, \varphi)_{\kappa} t = X t_{2R}$		
mm	millimeters.				
Y ₁ , Y ₂	The vectors of the image coordi- nates of left and right photos of a certain stereomodel.	$X_2 = \frac{1}{X_{2i}}$ $V = V_i$	2 8		
A_{i}	The matrix of partial derivatives of the image coordinates with respect to pass-point coordinates	$v = \overline{V}_2$ $W = F(L^b, X_1^o, X_2^o)$			

- The matrix of partial derivatives A_2 of the image coordinates with respect to orientation elements.
- X_1 The incremental correction for the pass point coordinates.
- X_{2} The incremental correction for the orientation elements.
- X_R The relative orientation elements $(\omega'', \varphi'', \varphi'', \kappa'', \kappa'').$
- σ 2 A priori variance of unit weight for the image coordinates.
- A priori variance of unit weight σ_{01}^2 related to the wideangle photog-

 $\mathbf{x}_1^a = \mathbf{X}_1 + \mathbf{X}_1^o$ $x_2^a = X_2 + X_2^o$ $Qy = \sigma_0^2$, Qy $P = Qy^{-1}$ $B = \frac{B_1}{B_2}$ $B_1 = {}^{\flat} \Upsilon 1 / {}^{\flat} M$ $B_2 = \lambda Y 2 / \lambda M$ $G_1 = \frac{3}{YI}/\frac{3}{XR}$ $G_2 = \frac{3}{2} \frac{7}{2} \frac{3}{2} R$ $\hat{\sigma}_o^2 = V^t P V/df$